The degeneracy of the effective activation barriers is lifted at high tunneling currents, and based on the Arrhenius-like switching behavior, the mean lifetimes $\tau_{0,1}$ can be expressed as (21, 22)

$$\tau_{0,1} = \frac{1}{v_0} \exp \left( \frac{E_b}{k_B T} \left( 1 - \frac{I}{I_c} \right) \right)$$

(1)

where $v_0$ is the attempt frequency, $E_b$ is the effective activation barrier of the island at zero current, $k_B$ is the Boltzmann constant, $I$ denotes the tunneling current, and $I_c$ is the threshold current to switch the magnetization at $T = 0$ K. $E_b$ was determined by a variation of $T$ from 50.6 to 48.5 K and derivation of the respective mean lifetimes $\tau_{0,1}$ at low tunneling currents ($I = 1$ nA). For the particular island of Fig. 3, we find $E_b = 133 \pm 4$ meV, leading to a threshold current of $I_c = 89 \pm 4$ µA. Because such high currents are not realizable within the tunneling regime, we do not expect to switch islands of the given dimensions at $T = 0$ K. Assuming an effective tunneling area given by the lateral STM resolution, the corresponding threshold current density is $(113 \pm 5) \times 10^8$ A/cm². This value is, by two to three orders of magnitude, higher than the current density used in similar experiments based on TMR devices (23), which may be attributed to the fact that, in contrast to planar junctions, the current density is not distributed homogeneously on the whole nanoisland but acts very locally. The splitting of the effective activation barrier $\Delta E$ due to spin-torque effects can be quantified by

$$\Delta E = k_B T \ln \left( \frac{1 + a_c}{1 - a_c} \right)$$

(2)

where $a_c$ is the lifetime asymmetry. For $I = 800$ nA, the current-induced spin torque leads to an effective activation barrier splitting of $\Delta E = 1.3 \pm 0.1$ meV, which is only $\approx 1\%$ of $E_b$.

Using a SP-STM tip as the source or drain for spin-polarized electrons, we were able to perform spatially resolved measurements where the tip is moved to different sites of one particular nanoisland, allowing information of site-specific properties to be gained that cannot be obtained in spatially averaging experiments performed with nanopillars. Figure 4A shows the topography of a nanoisland consisting of about 100 atoms. While scanning this island with $I = 600$ nA, we measured the magnetic $dI/dU$ signal on each of the pixels for a duration of 12 s to calculate the site-specific histogram asymmetry $a_{10}$ on the basis of the corresponding datapoint histograms. The result is shown in a color-coded representation in Fig. 4B. In spite of the rather large statistical error, a gradient along the [001] direction can clearly be recognized. The effect can even be analyzed quantitatively by averaging $a_{10}$ column- and row-wise: that is, along the [1¯10] and the [001] directions, respectively. Whereas $a_{10}$ is constant within the error at about 42% when moving the tip along the [1¯10] direction, it clearly reduces by about 16% from the left to the right side of the island. The lateral tip position obviously may influence the switching behavior of the nanoisland at high tunneling currents, as illustrated in Fig. 4, C to E. If the tip is positioned above the island center, the influence of the Oersted field along the [1¯10] direction, which is the easy axis of the nanoisland, cancels (Fig. 4C).

In this case, pure spin-current–induced switching occurs. Likewise, no influence by the Oersted field is expected if the tip is moved from the center to either island edge along the [1¯10] direction (Fig. 4D) as the effective field acting on the island is oriented perpendicular to the easy axis. Only if the tip is moved from the center along the [001] direction one magnetic state is favored over the other by the Oersted field (Fig. 4E). In this case, Oersted field effects influence the magnetic switching behavior, dependent on the tip position. A detailed analysis of the data yields that the effective activation barrier splitting of $\Delta E = -2.4 \pm 0.2$ meV at the center of the island is increased or decreased by up to $0.7 \pm 0.2$ meV when moving along the [001] direction to either island edge. This finding indicates that the magnetization switching is dominated by the spin torque induced by the spin-polarized current, whereas the influence of the Oersted field remains small. A simple estimation of the energy splitting caused by the Oersted field, as obtained by integrating the expected field distribution of an infinitely expanded current line over the island area, yields a Zeeman energy of up to $\Delta E_{\text{Oersted}} = \sum \mathbf{m} \cdot \mathbf{B} (r_i) \approx 0.2$ meV [$(m_r = 2.79 \mu_B (24)$], where m is the magnetic moment, $\mathbf{B}$ is the Oersted field, $r_i$ is the position of atom i, and $\mu_B$ is the Bohr magneton. This estimated value is somewhat lower than the experimental one. We attribute the difference to the over-simplified geometry of our model.

Our SP-STM studies provide insight into the details of current-induced magnetization switching that has been inaccessible in experiments with lithographically fabricated tunnel junctions. The ultimate lateral resolution of SP-STM combined with CIMS promises innovative perspectives for future data storage technologies.

**References and Notes**

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**Supporting Online Material**

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**SOM Text**

**Fig. S1**

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**Global Pattern Formation and Ethnic/Cultural Violence**

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We identify a process of global pattern formation that causes regions to differentiate by culture. Violence arises at boundaries between regions that are not sufficiently well defined. We model cultural differentiation as a separation of groups whose members prefer similar neighbors, with a characteristic group size at which violence occurs. Application of this model to the area of the former Yugoslavia and to India accurately predicts the locations of reported conflicts. This model also points to imposed mixing or boundary clarification as mechanisms for promoting peace.

Over the past 100 years, more than 100 million people have died in violent conflicts (1). Of these deaths, a great number are attributable to ongoing local conflict between culturally or ethnically distinct groups. A scientific understanding of the underlying causes of ethnic violence could lead to policy changes that may help stop or prevent it. The existing literature (2–13) [see also bibliography of ethnic and cultural conflict in the supporting online materials (14)] generally considers (i) the process by which ethnogenic identity is established and (if inter-
ventions could diminish its importance relative to more inclusive identities, and (ii) control mechanisms of the state and of organizations of ethnic groups and if interventions could strengthen the state while subsuming or accommodating ethnic groups within state authority. More specific social and economic factors identified in the literature as contributing to violence include oppression of minorities, economic grievances, historical precedents, competition for resources, favoritism, availability of resources for violence, security fears, mobilization by elites, weak social ties, national ethnic diversity, territorial claims, religious or political polarization, incendiary media, and international influences. Although most of these studies consider national conditions, a few consider local violence to identify the role of local socioeconomic or geographic factors (7–9). Here, we focus on an aspect of spatial population structure that has been neglected so far; we analyze the global pattern of violence and propose that many instances are consistent with the natural dynamics of type separation (15–18), a form of pattern formation (19) also seen in physical or chemical phase separation. Violence arises due to the structure of boundaries between groups rather than as a result of inherent conflicts between the groups themselves. In this approach, diverse social and economic causal factors trigger violence when the spatial population structure creates a propensity to conflict, so that spatial heterogeneity itself is predictive of local violence. The local ethnic patch size serves as an “order parameter,” a measure of the degree of order of collective behavior, to which other aspects of behavior are coupled. The importance of collective behavior implies that ethnic violence can be studied in the universal context of collective dynamics, where models can identify how individual and collective behavior are related.

A simple model of type separation is shown in Fig. 1, A to E. The dynamics of this model assume that individuals preferentially move to areas where more individuals of the same type reside (14). The resulting dynamics lead to progressively larger patches (“islands” or “peninsulas”) of each type. The average size of patches at a particular time can be obtained by a number of different methods. We used overlapping spatial waves that represent the spatial variation of the population density. Each wave makes a contribution proportional to its correlation with the population density (the structure factor or Fourier transform). The wavelength of the wave that has the maximum amplitude gives the average size of the patches. Other methods of obtaining the size of patches give similar results. The size of the patches grows as a characteristic power of time (Fig. 1F, inset). This behavior has been proven (20) to be a “universal behavior” that does not depend on many of the details of the model and therefore may be relied on to describe a large variety of systems of interacting elements; in particular, similar models have been used to describe the relation of chemical interaction energies and chemical precipitation or phase separation (21, 22). The universal properties of the patterns upon rescaling of length and time also imply that a number of individual agents of the model can be aggregated into a single agent if time is rescaled correspondingly without changing the behavior at the larger scales (Fig. 1F). Thus, it is possible to consider a model agent to represent a local population, and it is not necessary to model the behavior of each individual—an impractical undertaking.

To model violence, we assume that highly mixed regions do not engage in violence, and neither do well-segregated groups, an intuitive hypothesis with empirical support (7). The analy-

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**Fig. 1.** Simulation of type separation with two types of agents [(A) to (E) show the system at 8, 64, 512, 4096, and 32768 attempted moves per particle, respectively]. The shape of domains (as characterized by the rescaled structure factor amplitude squared) remains constant after an initial transient (F), and the average size of clusters grows as a power law [inset of (F)] (14). Patches of a certain size that are surrounded by the other type are highlighted by red shading overlay in (A) to (E). We identify such regions with a high likelihood of conflict.
sis is applicable to communal violence and not to criminal activity or interstate warfare. In highly mixed regions, groups of the same type are not large enough to develop strong collective identities, or to identify public spaces as associated with one or another cultural group. They are neither imposed upon nor impose upon other groups, and are not perceived as a threat to the cultural values or social/political self-determination of other groups. Partial separation with poorly defined boundaries fosters conflict. Violence arises when groups are of a size that they are able to impose cultural norms on public spaces, but where there are still intermittent violations of these rules due to the overlap of cultural domains. When groups are larger than the critical size, they typically form self-sufficient entities that enjoy local sovereignty. Hence, we expect violence to arise when groups of a certain characteristic size are formed, and not when groups are much smaller or larger than this size. The model of violence depends on the distribution of the population and not on the specific mechanism by which the population achieves this structure, which may include internally or externally directed migrations. By focusing on the geographic distribution of the population, the model seeks a predictor of conflict that can be easily determined by census. This may work well because geography is an important aspect of the dimensions of social space, the dynamic coarsening process is universal, and other aspects of social behavior (e.g., isolationism, conformity, as well as violence) are correlated to it.

The predictor that we identify based on spatial census data need not describe the immediate social or institutional triggers of violence, only the conditions under which violence becomes likely. Previous research aiming to characterize ethnic conflict by census data has focused on measures of ethnic or religious “fragmentation” (23–27). Such measures characterize the diversity of a country without reference to its spatial structure, i.e., the overall proportions of ethnically distinct groups in a country. They are therefore distinct from the spatial characterization of our study. The literature is divided about whether or which correlations exist with measures of national ethnic composition. We

![Fig. 2.](image-url)

**Fig. 2.** (A) Census data from 1991 shown here in map form were converted into a spatial representation and used in an agent-based simulation shown in (B). Our prediction of populations likely to be in conflict with neighboring groups [red overlay, (C) and (D)] agrees well with the location of cities reported as sites of major fights and massacres [yellow dots, (D)].
find, however, that the spatial distribution of ethnic groups is a strong predictor of locations of violence.

Mathematically, the expected violence was determined by detecting patches consisting of islands or peninsulas of one type surrounded by populations of other types. We detected these features by correlation of the population for each population type with a template that has a positive center and a negative surround. To illustrate the effect of this correlation, for a particular template size, the maximum correlation over population types is superimposed as a red overlay in Fig. 1, A to E. Over time in this simulation, the patch size starts smaller, then passes through and becomes larger than the template size chosen. The specific template that we used is based on a wavelet filter (14, 28–30). Wavelets are designed to obtain a local measure of the degree to which a certain scale of variation (wavelength) is present. Outcomes are highly robust, and other templates give similar results. Given the universality of the dynamic behavior, the diameter of the positive region of the wavelet, i.e., the size of the local population patches that are likely to experience violence, is the only essential parameter of the model. The parameter is to be determined by agreement of the model with reports of violence, though as we will see, the agreement is robust to variation of the parameter. The quality of the agreement provides a measure of the validity of the model.

To test the predictive ability of the model, we performed simulations based on census data for the former Yugoslavia and India. We assigned areas of pixelated geographic maps pixel by pixel to ethnic groups at random, but in proportion to their relative population census in the region. Although this does not reflect the physical geography or local mixing of groups in buildings and villages, over an area of multiple pixels it captures the regional composition of the census. The pixelated map serves as the beginning state for the agent model. For Yugoslavia, census data from the early 1990s before the outbreak of conflict (31, 32), as shown in Fig. 2A, were captured into an agent simulation (Fig. 2B), which was used to obtain the regions of expected violence shown in Fig. 2C.

We then obtained from books (2), newspapers, and Internet sources (see supporting online text) the locations of reported violence for the area of the former Yugoslavia. Multiple independent sources were used to provide validation for each location of violence (14). We consider these reports as indicators of areas of actual violence, keeping in mind possible bias and incompleteness and that areas of widespread violence are identified only by local urban centers. In comparing such reports with model predictions, we note that the model identifies locations of groups of a particular size, but the location of the actual violence should occur somewhere in the area between adjacent groups. Despite these caveats, overlaying the locations of reported and predicted violence in Fig. 2D demonstrates a significant ability of our simple model to identify regions of reported violence. We performed statistical analyses comparing the predicted to the reported violence, evaluating the ability of the model to determine both where violence occurs and where violence does not occur. For comparison, we randomized the locations of reported violence. We defined “conflict proximity” as the distance between a given position and the nearest location of violence (predicted, reported, or randomized). We calculated Pearson’s correlation and other statistical measures between the proximities of predicted and reported violence, and compared them with the same measures in relation to randomized reports. We found that the model has a correlation of 0.9 with reports (0.89 to two significant digits), a level of agreement not reached in any of 100,000 randomized trials. Moreover, the predicted results are highly robust to parameter variation, with essentially equivalent agreement obtained for filter diameters ranging from 18 to 60 km, a range that is in agreement with intuition about the size of conflict areas. Below or above this range, poorer agreement occurs. Details are provided in the supporting online text.

We studied conflict in India as a second case study of the ethnic violence model. We constructed a spatial representation of India on a district level from maps at www.censusindia.net and obtained the distribution of ethno-cultural groups from the 2001 Census data at www.indiastat.com. The result can be seen in the form of three-color maps in Fig. 3, A and B, representing the relative densities of Hindus, Muslims, Christians, Sikhs, Buddhists, and Others (primarily Jains). The agent model is shown in Fig. 3C and the prediction of ethnic violence is indicated in Fig. 3D. Predictions correspond very well to the primary locations of “extremist” violence of government reports as given by indiastat.com (Fig. 3E) and confirmed by independent sources (14), particularly in Kashmir, Punjab, and the states of Northeast India. Some additional areas of lesser violence were also predicted by the model, particularly Jharkhand—an eastern state created in 2000 that has recently experienced some violence (14, 33). Consistent with predicted results, the violence in this region is not as prevalent as in other violence-prone areas of India. Statistical correlation measures of conflict proximity yield a correlation of 0.998 when the threshold is set above the value of predicted violence in Jharkhand. If the threshold is set lower, so that violence in Jharkhand is included in predicted but not in reported cases, the correlation falls to 0.92. Including reported violence in Jharkhand when comparing at the lower threshold increases the correlation to 0.98. Additional details are provided in the supporting online text. The range of filter diameter values for which good agreement was obtained overlaps that of the former Yugoslavia. However, it is shifted to larger values, up to ~100 km. This may reflect not only the larger granularity of data, but perhaps also the effect of violence itself on separation. Unlike Yugoslavia, in India the census was performed during ongoing violence. Because violence accelerates the process of separation, groups in conflict are likely to have separated substantially and reflect the high end of group sizes susceptible to violence.

Governmental and nongovernmental organizations are devoting increasing attention to the prevention of major conflict (34). Under some circumstances, social and institutional factors that affect violence might serve to suppress the triggering of violence without changing the spatial structure of the population. However, influencing the spatial structure might address the conditions that promote violence described here. Such ap-
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14. Materials and methods are available as supporting material on Science Online.

Crystal Structure of an Ancient Protein: Evolution by Conformational Epistasis

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The structural mechanisms by which proteins have evolved new functions are known only indirectly. We report x-ray crystal structures of a resurrected ancestral protein—the ~450 million-year-old precursor of vertebrate glucocorticoid (GR) and mineralocorticoid (MR) receptors. Using structural, phylogenetic, and functional analysis, we identify the specific set of historical mutations that recapitulate the evolution of GR’s hormone specificity from an MR-like ancestor. These substitutions repositioned crucial residues to create new receptor-ligand and intraprotein contacts. Strong epistatic interactions occur because one substitution changes the conformational position of another site. “Permissive” mutations—substitutions of no immediate consequence, which stabilize specific elements of the protein and allow it to tolerate subsequent function-switching changes—played a major role in determining GR’s evolutionary trajectory.

A central goal in molecular evolution is to understand the mechanisms and dynamics by which changes in gene sequence generate shifts in function and therefore phenotype (1, 2). A complete understanding of this process requires analysis of how changes in protein structure mediate the effects of mutations on function. Comparative analyses of extant proteins have provided indirect insights into the diversification of protein structure (3–6), and protein